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# Supplementary Material for “Extending multivariate- $t$ linear mixed models for multiple longitudinal data with censored responses and heavy tails”

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This supporting information is a longer version of the printed paper. It contains the detailed proofs of Propositions 2 and 3, and the missing information matrix of  $\beta$  together with Supplementary Table for the simulation study.

## Web Appendix A: Proof of Proposition 2.

Using the fact of  $\mathbf{y}_i = \mathbf{O}_i^T \mathbf{y}_i^o + \mathbf{C}_i^T \mathbf{y}_i^c$ , it follows from Proposition 3 of Matos et al. (2013) and Proposition 1 of this paper that

$$\begin{aligned}\widehat{\tau \mathbf{y}}_i^{(h)} &= E[\tau_i \mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i | \mathbf{y}_i] (\mathbf{O}_i^T \mathbf{y}_i^o + \mathbf{C}_i^T \mathbf{y}_i^c) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= \mathbf{O}_i^T E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \mathbf{y}_i^o + \mathbf{C}_i^T E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \hat{\tau}_i^{(h)} (\mathbf{O}_i^T \mathbf{y}_i^o + \mathbf{C}_i^T E[\mathbf{W}_i^c]),\end{aligned}$$

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and

$$\begin{aligned}
\widehat{\tau \mathbf{y}_i^2}^{(h)} &= E[\tau_i \mathbf{y}_i \mathbf{y}_i^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i | \mathbf{y}_i] (\mathbf{O}_i^T \mathbf{y}_i^o + \mathbf{C}_i^T \mathbf{y}_i^c) (\mathbf{O}_i^T \mathbf{y}_i^o + \mathbf{C}_i^T \mathbf{y}_i^c)^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \mathbf{O}_i^T \mathbf{y}_i^o \mathbf{y}_i^{oT} \mathbf{O}_i + \mathbf{O}_i^T \mathbf{y}_i^o E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c{}^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \mathbf{C}_i \\
&\quad + \mathbf{C}_i^T E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \mathbf{y}_i^{oT} \mathbf{O}_i + \mathbf{C}_i^T E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c \mathbf{y}_i^c{}^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \mathbf{C}_i \\
&= \hat{\tau}_i^{(h)} \left\{ \mathbf{O}_i^T \mathbf{y}_i^o \mathbf{y}_i^{oT} \mathbf{O}_i + \mathbf{O}_i^T \mathbf{y}_i^o E^T[\mathbf{W}_i^c] \mathbf{C}_i + \mathbf{C}_i^T E[\mathbf{W}_i^c] \mathbf{y}_i^{oT} \mathbf{O}_i + \mathbf{C}_i^T E[\mathbf{W}_i^c \mathbf{W}_i^c{}^T] \mathbf{C}_i \right\},
\end{aligned}$$

where the detailed derivations of  $\hat{\tau}_i^{(h)}$  and  $\mathbf{W}_i^c$  are given in the proof of Web Appendix B. Note that throughout the supplementary material, we use the notation, e.g.,  $E_X[f(X, \dots) | y, z, \dots]$ , to indicate that the term inside  $f(X, \dots)$  needed to be integrated over the conditional distribution  $X$  given  $y, z, \dots$  is only  $X$  under this conditional expectation.

According to the hierarchical formulation of MtLMMC given in (4), we use the standard matrix factorizations (cf. Anderson (2003), Appendix A.3) to obtain

$$\mathbf{b}_i | \mathbf{y}_i, \tau_i \sim \mathcal{N}_q(\boldsymbol{\Sigma}_{\mathbf{b}_i} \mathbf{Z}_i^T \mathbf{R}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}), \tau_i^{-1} \boldsymbol{\Sigma}_{\mathbf{b}_i}).$$

This conditional distribution is useful for evaluating the desired conditional moments of latent data, including

$$\begin{aligned}
\widehat{\tau \mathbf{b}_i}^{(h)} &= E[\tau_i \mathbf{b}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i E_{\mathbf{b}_i} (\mathbf{b}_i | \mathbf{y}_i, \tau_i) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i \boldsymbol{\Sigma}_{\mathbf{b}_i} \mathbf{Z}_i^T \mathbf{R}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [\boldsymbol{\Sigma}_{\mathbf{b}_i} \mathbf{Z}_i^T \mathbf{R}_i^{-1} E_{\tau_i} [\tau_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} \left[ \boldsymbol{\Sigma}_{\mathbf{b}_i} \mathbf{Z}_i^T \mathbf{R}_i^{-1} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) \right] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\
&= \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \mathbf{Z}_i^T \hat{\mathbf{R}}_i^{(h)-1} (\widehat{\tau \mathbf{y}_i}^{(h)} - \hat{\tau}_i^{(h)} \mathbf{X}_i \hat{\boldsymbol{\beta}}^{(h)}),
\end{aligned}$$

$$\begin{aligned}
\widehat{\tau \mathbf{b}_i^2}^{(h)} &= E[\tau_i \mathbf{b}_i \mathbf{b}_i^\top | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i E_{\mathbf{b}_i} (\mathbf{b}_i \mathbf{b}_i^\top | \mathbf{y}_i, \tau_i) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i (\boldsymbol{\Sigma}_{\mathbf{b}_i} \mathbf{Z}_i^\top \mathbf{R}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})^\top \mathbf{R}_i^{-1} \mathbf{Z}_i \boldsymbol{\Sigma}_{\mathbf{b}_i} + \tau_i^{-1} \boldsymbol{\Sigma}_{\mathbf{b}_i}) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \mathbf{Z}_i^\top \hat{\mathbf{R}}_i^{(h)-1} E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})^\top | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \hat{\mathbf{R}}_i^{(h)-1} \mathbf{Z}_i \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} + \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \\
&= \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \mathbf{Z}_i^\top \hat{\mathbf{R}}_i^{(h)-1} E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) (\mathbf{y}_i \mathbf{y}_i^\top - \mathbf{y}_i \boldsymbol{\beta}^\top \mathbf{X}_i^\top - \mathbf{X}_i \boldsymbol{\beta} \mathbf{y}_i^\top + \mathbf{X}_i \boldsymbol{\beta} \boldsymbol{\beta}^\top \mathbf{X}_i^\top) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\
&\quad \times \hat{\mathbf{R}}_i^{(h)-1} \mathbf{Z}_i \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} + \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \\
&= \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \mathbf{Z}_i^\top \hat{\mathbf{R}}_i^{(h)-1} \left[ \widehat{\tau \mathbf{y}_i^2}^{(h)} - \widehat{\tau \mathbf{y}_i}^{(h)} \hat{\boldsymbol{\beta}}^{(h)\top} \mathbf{X}_i^\top - \mathbf{X}_i \boldsymbol{\beta}^{(h)} \widehat{\tau \mathbf{y}_i}^{(h)\top} + \hat{\tau}_i^{(h)} \mathbf{X}_i \hat{\boldsymbol{\beta}}^{(h)} \hat{\boldsymbol{\beta}}^{(h)\top} \mathbf{X}_i^\top \right] \\
&\quad \times \hat{\mathbf{R}}_i^{(h)-1} \mathbf{Z}_i \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} + \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)},
\end{aligned}$$

$$\begin{aligned}
\widehat{\tau \mathbf{y}_i \mathbf{b}_i}^{(h)} &= E[\tau_i \mathbf{y}_i \mathbf{b}_i^\top | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i \mathbf{y}_i E_{\mathbf{b}_i} (\mathbf{b}_i^\top | \mathbf{y}_i, \tau_i) | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i \mathbf{y}_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})^\top \mathbf{R}_i^{-1} \mathbf{Z}_i \boldsymbol{\Sigma}_{\mathbf{b}_i} | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) [\mathbf{y}_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})^\top | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \hat{\mathbf{R}}_i^{(h)-1} \mathbf{Z}_i \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)} \right] \\
&= (\widehat{\tau \mathbf{y}_i^2}^{(h)} - \widehat{\tau \mathbf{y}_i}^{(h)} \hat{\boldsymbol{\beta}}^{(h)\top} \mathbf{X}_i^\top) \hat{\mathbf{R}}_i^{(h)-1} \mathbf{Z}_i \hat{\boldsymbol{\Sigma}}_{\mathbf{b}_i}^{(h)},
\end{aligned}$$

and

$$\begin{aligned}
\widehat{\tau \mathbf{E}_i}^{(h)} &= E[\tau_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i) (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i)^\top | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i E_{\mathbf{b}_i} ((\mathbf{y}_i \mathbf{y}_i^\top - \mathbf{y}_i \boldsymbol{\beta}^\top \mathbf{X}_i^\top - \mathbf{y}_i \mathbf{b}_i^\top \mathbf{Z}_i^\top - \mathbf{X}_i \boldsymbol{\beta} \mathbf{y}_i^\top - \mathbf{Z}_i \mathbf{b}_i \mathbf{y}_i^\top + \mathbf{X}_i \boldsymbol{\beta} \boldsymbol{\beta}^\top \mathbf{X}_i^\top \\
&\quad + \mathbf{Z}_i \mathbf{b}_i \mathbf{b}_i^\top \mathbf{Z}_i^\top + \mathbf{X}_i \boldsymbol{\beta} \mathbf{b}_i^\top \mathbf{Z}_i^\top + \mathbf{Z}_i \mathbf{b}_i \boldsymbol{\beta}^\top \mathbf{X}_i^\top) | \mathbf{y}_i, \tau_i) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\
&= \widehat{\tau \mathbf{y}_i^2}^{(h)} - \mathbf{X}_i \boldsymbol{\beta} (\widehat{\tau \mathbf{y}_i}^{(h)\top} - \widehat{\tau \mathbf{b}_i}^{(h)\top} \mathbf{Z}_i^\top) - (\widehat{\tau \mathbf{y}_i}^{(h)} - \mathbf{Z}_i \widehat{\tau \mathbf{b}_i}^{(h)}) \boldsymbol{\beta}^\top \mathbf{X}_i \\
&\quad - \widehat{\tau \mathbf{y}_i \mathbf{b}_i}^{(h)} \mathbf{Z}_i^\top - \mathbf{Z}_i \widehat{\tau \mathbf{y}_i \mathbf{b}_i}^{(h)} + \hat{\tau}_i^{(h)} \mathbf{X}_i \boldsymbol{\beta} \boldsymbol{\beta}^\top \mathbf{X}_i^\top + \mathbf{Z}_i \widehat{\tau \mathbf{b}_i^2}^{(h)} \mathbf{Z}_i^\top.
\end{aligned}$$

This completes the proof of Proposition 2.  $\square$

## Web Appendix B: Proof of Proposition 3.

From hierarchy (6), dividing the joint PDF of  $(\mathbf{y}_i, \tau_i)$  by the marginal PDF of  $\mathbf{y}_i$  yields

$$\tau_i | \mathbf{y}_i \sim \text{Gamma} \left( \frac{\nu + n_i}{2}, \frac{\nu + \Delta_{\mathbf{y}_i}}{2} \right),$$

where  $\Delta_{\mathbf{y}_i} = (\mathbf{y}_i - \mathbf{X}_i\boldsymbol{\beta})^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i\boldsymbol{\beta})$ . Using this fact, we can derive the formulae for  $\hat{\tau}_i^{(h)}$  and the distribution of  $\mathbf{W}_i^c$  under the following three censoring patterns, say (i) only non-censored, (ii) only censored, and (iii) both censored and non-censored situations.

(i) When the  $i$ th subject has only non-censored (observed) measurements, we have  $\mathbf{u}_i = \mathbf{y}_i$ , and thus

$$\begin{aligned}\hat{\tau}_i^{(h)} &= E[\tau_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] = E[\tau_i | \mathbf{y}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] = (\hat{\nu}^{(h)} + n_i) / (\hat{\nu}^{(h)} + \hat{\Delta}_{\mathbf{y}_i}^{(h)}), \\ \widehat{\tau\mathbf{y}}_i^{(h)} &= E[\tau_i \mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] = E_{\tau_i} [\tau_i | \mathbf{y}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \mathbf{y}_i = \hat{\tau}_i^{(h)} \mathbf{y}_i, \\ \widehat{\tau\mathbf{y}}_i^2{}^{(h)} &= E[\tau_i \mathbf{y}_i \mathbf{y}_i^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] = E_{\tau_i} [\tau_i | \mathbf{y}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \mathbf{y}_i \mathbf{y}_i^T = \hat{\tau}_i^{(h)} \mathbf{y}_i \mathbf{y}_i^T.\end{aligned}$$

(ii) When the  $i$ th subject has only censored measurements, the auxiliary permutation matrix  $\mathbf{O}_i$  is a null set and  $\mathbf{C}_i = \mathbf{I}_{n_i}$ . Thus  $\mathbf{y}_i = \mathbf{y}_i^c$ , and  $\mathbf{y}_i \sim \mathcal{T}t_{n_i}(\mathbf{X}_i\boldsymbol{\beta}, \frac{\nu}{\nu+2}\boldsymbol{\Lambda}_i, \nu+2; \mathbb{A}_i)$ , where  $\mathbb{A}_i = \{\mathbf{y}_i = [y_{ijk}] \mid y_{ijk} \leq u_{ijk}, j = 1, \dots, r, k = 1, \dots, s_i\}$ . Accordingly,

$$\begin{aligned}\hat{\tau}_i^{(h)} &= E[\tau_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i | \mathbf{y}_i] | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}^{(h)}} \right) | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \frac{\mathcal{T}_{n_i}(\mathbf{u}_i | \mathbf{X}_i \hat{\boldsymbol{\beta}}^{(h)}, \frac{\hat{\nu}^{(h)}}{\hat{\nu}^{(h)}+2} \hat{\boldsymbol{\Lambda}}_i^{(h)}, \hat{\nu}^{(h)} + 2)}{\mathcal{T}_{n_i}(\mathbf{u}_i | \mathbf{X}_i \hat{\boldsymbol{\beta}}^{(h)}, \hat{\boldsymbol{\Lambda}}_i^{(h)}, \hat{\nu}^{(h)})},\end{aligned}\tag{S.1}$$

in which the last equality follows from Proposition 2 of Matos et al. (2013) with  $k = 0$ . Furthermore, by Proposition 2 of Matos et al. (2013) with  $k = 1$  and 2, we obtain

$$\begin{aligned}\widehat{\tau\mathbf{y}}_i^{(h)} &= E[\tau_i \mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i | \mathbf{y}_i] \mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}^{(h)}} \right) \mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \hat{\tau}_i^{(h)} E[\mathbf{W}_i^c],\end{aligned}$$

and

$$\begin{aligned}\widehat{\tau\mathbf{y}}_i^2{}^{(h)} &= E[\tau_i \mathbf{y}_i \mathbf{y}_i^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} [E_{\tau_i} [\tau_i | \mathbf{y}_i] \mathbf{y}_i \mathbf{y}_i^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}^{(h)}} \right) \mathbf{y}_i \mathbf{y}_i^T | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \hat{\tau}_i^{(h)} E[\mathbf{W}_i^c \mathbf{W}_i^c{}^T],\end{aligned}$$

where  $\hat{\tau}_i^{(h)}$  is given by (S.1), and  $\mathbf{W}_i^c \sim \mathcal{T}t_{n_i}(\mathbf{X}_i \hat{\boldsymbol{\beta}}^{(h)}, \frac{\hat{\nu}^{(h)}}{\hat{\nu}^{(h)}+2} \hat{\boldsymbol{\Lambda}}_i^{(h)}, \hat{\nu}^{(h)} + 2; \mathbb{A}_i)$  with  $\hat{\boldsymbol{\Lambda}}_i^{(h)}$  being  $\boldsymbol{\Lambda}_i$  evaluated at  $\boldsymbol{\theta} = \hat{\boldsymbol{\theta}}^{(h)}$ .

(iii) When the  $i$ th subject has both censored and non-censored measurements, using the fact that  $\{\mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i\}$ ,  $\{\mathbf{y}_i | \mathbf{u}_i, \mathbf{c}_i, \mathbf{y}_i^o\}$ , and  $\{\mathbf{y}_i^c | \mathbf{u}_i, \mathbf{c}_i, \mathbf{y}_i^o\}$  are equivalent processes, we have

$$\begin{aligned} \hat{\tau}_i^{(h)} &= E[\tau_i | \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i^c} [E_{\tau_i} [\tau_i | \mathbf{y}_i] | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i^c} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \left( \frac{\hat{\nu}^{(h)} + n_i^o}{\hat{\nu}^{(h)} + \hat{\Delta}_{\mathbf{y}_i^o}^{(h)}} \right) \frac{\mathcal{T}_{n_i - n_i^o}(\mathbf{u}_i^c | \hat{\boldsymbol{\mu}}_i^{c \cdot o^{(h)}}, \frac{\hat{\nu}^{(h)} + n_i^o}{\hat{\nu}^{(h)} + n_i^o + 2} \hat{\mathbf{S}}_i^{cc \cdot o^{(h)}}, \hat{\nu}^{(h)} + n_i^o + 2)}{\mathcal{T}_{n_i - n_i^o}(\mathbf{u}_i^c | \hat{\boldsymbol{\mu}}_i^{c \cdot o^{(h)}}, \hat{\mathbf{S}}_i^{cc \cdot o^{(h)}}, \hat{\nu}^{(h)} + n_i^o)}, \quad (\text{S.2}) \end{aligned}$$

in which the last equality follows from the conditional distribution of  $\mathbf{y}_i^c | \mathbf{y}_i^o$  specified in Proposition 1 of this paper and Proposition 3 of Matos et al. (2013) with  $k = 0$ , where  $\hat{\boldsymbol{\Lambda}}_{\mathbf{y}_i^o}^{(h)} = (\mathbf{y}_i^o - \mathbf{X}_i^o \hat{\boldsymbol{\beta}}^{(h)})^T \hat{\boldsymbol{\Lambda}}_i^{oo^{(h)}}^{-1} (\mathbf{y}_i^o - \mathbf{X}_i^o \hat{\boldsymbol{\beta}}^{(h)})$  with  $\hat{\boldsymbol{\Lambda}}_i^{oo^{(h)}}$  being  $\boldsymbol{\Lambda}_i^{oo}$  evaluated at  $\boldsymbol{\theta} = \hat{\boldsymbol{\theta}}^{(h)}$ . Under this case, using Proposition 3 of Matos et al. (2013) again with  $k = 1$  and 2, the conditional moments involving the latent data  $\tau_i$  and  $\mathbf{y}_i^c$  within  $\widehat{\tau \mathbf{y}_i}^{(h)}$  and  $\widehat{\tau \mathbf{y}_i^2}^{(h)}$  in Proposition 2 become

$$\begin{aligned} E[\tau_i \mathbf{y}_i^c | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] &= E_{\mathbf{y}_i^c} [E_{\tau_i} [\tau_i | \mathbf{y}_i^c] \mathbf{y}_i^c | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i^c} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \hat{\tau}_i^{(h)} E[\mathbf{W}_i^c], \end{aligned}$$

and

$$\begin{aligned} E[\tau_i \mathbf{y}_i^c \mathbf{y}_i^{cT} | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] &= E_{\mathbf{y}_i^c} [E_{\tau_i} [\tau_i | \mathbf{y}_i] \mathbf{y}_i^c \mathbf{y}_i^{cT} | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)}] \\ &= E_{\mathbf{y}_i^c} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{y}_i^c \mathbf{y}_i^{cT} | \mathbf{y}_i^o, \mathbf{u}_i, \mathbf{c}_i, \hat{\boldsymbol{\theta}}^{(h)} \right] \\ &= \hat{\tau}_i^{(h)} E[\mathbf{W}_i^c \mathbf{W}_i^{cT}], \end{aligned}$$

where  $\hat{\tau}_i^{(h)}$  is given by (S.2), and  $\mathbf{W}_i^c \sim \mathcal{T}t_{n_i - n_i^o}(\hat{\boldsymbol{\mu}}_i^{c \cdot o^{(h)}}, \frac{\hat{\nu}^{(h)} + n_i^o}{\hat{\nu}^{(h)} + n_i^o + 2} \hat{\mathbf{S}}_i^{cc \cdot o^{(h)}}, \hat{\nu}^{(h)} + n_i^o + 2; \mathbb{A}_i^c)$  with  $\hat{\boldsymbol{\mu}}_i^{c \cdot o^{(h)}}$  and  $\hat{\mathbf{S}}_i^{cc \cdot o^{(h)}}$  being  $\boldsymbol{\mu}_i^{c \cdot o}$  and  $\mathbf{S}_i^{cc \cdot o}$ , respectively, evaluated at  $\boldsymbol{\theta} = \hat{\boldsymbol{\theta}}^{(h)}$ , and  $\mathbb{A}_i^c$  representing the left truncated region at  $[\mathbf{w}_i^c \leq \mathbf{u}_i^c]$ .

The desired quantities have been formulated. □

### Web Appendix C: Proof of the missing information matrix of $\beta$ .

The missing information matrix of  $\beta$  can be calculated as the conditional variance-covariance matrix of complete-data score vector with respect to  $\beta$ , that is,

$$\begin{aligned}
\mathbf{I}_m(\beta; \mathbf{y}) &= \sum_{i=1}^N \text{Var} \left( \frac{\partial \ell_c(\boldsymbol{\theta} | \text{Data})}{\partial \beta} \middle| \mathbf{u}_i, \mathbf{c}_i \right) \\
&= \sum_{i=1}^N \text{Var} \left( \tau_i \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{u}_i, \mathbf{c}_i \right) \\
&= \sum_{i=1}^N \left\{ \text{Var}_{\mathbf{y}_i} \left( E_{\tau_i} \left[ \tau_i \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{y}_i \right] \middle| \mathbf{u}_i, \mathbf{c}_i \right) \right. \\
&\quad \left. + E_{\mathbf{y}_i} \left( \text{Var}_{\tau_i} \left[ \tau_i \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{y}_i \right] \middle| \mathbf{u}_i, \mathbf{c}_i \right) \right\} \\
&= \sum_{i=1}^N \left\{ \text{Var}_{\mathbf{y}_i} \left( \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{u}_i, \mathbf{c}_i \right) \right. \\
&\quad \left. + E_{\mathbf{y}_i} \left( \left[ \frac{2(\nu + n_i)}{(\nu + \Delta_{\mathbf{y}_i})^2} \right] \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \middle| \mathbf{u}_i, \mathbf{c}_i \right) \right\} \\
&= \sum_{i=1}^N \left\{ E_{\mathbf{y}_i} \left( \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right)^2 \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \middle| \mathbf{u}_i, \mathbf{c}_i \right) \right. \\
&\quad - E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{u}_i, \mathbf{c}_i \right] E_{\mathbf{y}_i}^T \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{u}_i, \mathbf{c}_i \right] \\
&\quad \left. + E_{\mathbf{y}_i} \left[ \left( \frac{2}{\nu + n_i} \right) \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right)^2 \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \beta) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \middle| \mathbf{u}_i, \mathbf{c}_i \right] \right\} \\
&= \sum_{i=1}^N \left\{ \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} E_{\mathbf{y}_i} \left( \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right)^2 (\mathbf{y}_i - \mathbf{X}_i \beta) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \middle| \mathbf{u}_i, \mathbf{c}_i \right) \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \right. \\
&\quad - \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) (\mathbf{y}_i - \mathbf{X}_i \beta) \middle| \mathbf{u}_i, \mathbf{c}_i \right] E_{\mathbf{y}_i} \left[ \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \middle| \mathbf{u}_i, \mathbf{c}_i \right] \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \\
&\quad \left. + \left( \frac{2}{\nu + n_i} \right) \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} E_{\mathbf{y}_i} \left( \left( \frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}} \right)^2 (\mathbf{y}_i - \mathbf{X}_i \beta) (\mathbf{y}_i - \mathbf{X}_i \beta)^T \middle| \mathbf{u}_i, \mathbf{c}_i \right) \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \right\} \\
&= \sum_{i=1}^N \left\{ \mathbf{X}_i^T \boldsymbol{\Lambda}_i^{-1} \left[ \left( 1 + \frac{2}{\nu + n_i} \right) (\widehat{\tau^2 \mathbf{y}_i^2} - \widehat{\tau^2 \mathbf{y}_i} \beta^T \mathbf{X}_i^T - \mathbf{X}_i \beta \widehat{\tau^2 \mathbf{y}_i}^T + \widehat{\tau_i^2} \mathbf{X}_i \beta \beta^T \mathbf{X}_i^T) \right. \right. \\
&\quad \left. \left. - (\widehat{\tau \mathbf{y}_i} - \widehat{\tau_i} \mathbf{X}_i \beta) (\widehat{\tau \mathbf{y}_i} - \widehat{\tau_i} \mathbf{X}_i \beta)^T \right] \boldsymbol{\Lambda}_i^{-1} \mathbf{X}_i \right\}
\end{aligned}$$

where  $\widehat{\tau}\mathbf{y}_i$  and  $\widehat{\tau}_i$  are  $\widehat{\tau}\mathbf{y}_i^{(h)}$  and  $\widehat{\tau}_i^{(h)}$  given in Propositions 2 and 3, respectively, with  $\widehat{\boldsymbol{\theta}}^{(h)}$  replaced by  $\boldsymbol{\theta}$ ; and the proof of expressions for  $\widehat{\tau}_i^{(h)}$ ,  $\widehat{\tau}^2\mathbf{y}_i^{(h)}$ , and  $\widehat{\tau}^2\mathbf{y}_i^{(h)2}$  are derived as follows under three censoring situations:

1. If the  $i$ th subject has only non-censored measurements, then it is straightforward that

$$\widehat{\tau}_i^2 = \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2, \quad \widehat{\tau}\mathbf{y}_i = \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i, \quad \text{and} \quad \widehat{\tau}^2\mathbf{y}_i^2 = \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \mathbf{y}_i^T.$$

2. If the  $i$ th subject has only censored measurements, then we use the property of Proposition 2 of Matos et al. (2013) to derive the following conditional expectation:

$$\begin{aligned} \widehat{\tau}_i^2 &= E_{\mathbf{y}_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \middle| \mathbf{u}_i, \mathbf{c}_i \right] = c_{n_i}(\nu, 2) \frac{\mathcal{T}_{n_i}(\mathbf{u}_i | \mathbf{X}_i \boldsymbol{\beta}, \frac{\nu}{\nu+4} \boldsymbol{\Lambda}_i, \nu + 4)}{\mathcal{T}_{n_i}(\mathbf{u}_i | \mathbf{X}_i \boldsymbol{\beta}, \boldsymbol{\Lambda}_i, \nu)}, \\ \widehat{\tau}^2\mathbf{y}_i &= E_{\mathbf{y}_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \middle| \mathbf{u}_i, \mathbf{c}_i \right] = \widehat{\tau}_i^2 E[\mathbf{W}_i^*], \end{aligned}$$

and

$$\widehat{\tau}^2\mathbf{y}_i^2 = E_{\mathbf{y}_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \mathbf{y}_i^T \middle| \mathbf{u}_i, \mathbf{c}_i \right] = \widehat{\tau}_i^2 E(\mathbf{W}_i^* \mathbf{W}_i^{*T}),$$

where  $\mathbf{W}_i^* \sim \mathcal{T}t_{n_i}(\mathbf{X}_i \boldsymbol{\beta}, \frac{\nu}{\nu+4} \boldsymbol{\Lambda}_i, \nu + 4, \mathbb{A}_i)$ , and

$$c_{n_i}(\nu, 2) = \left(\frac{\nu + n_i}{\nu}\right)^2 \frac{\Gamma(\frac{n_i + \nu}{2}) \Gamma(\frac{\nu + 4}{2})}{\Gamma(\frac{\nu}{2}) \Gamma(\frac{n_i + \nu + 4}{2})} = \frac{(\nu + n_i)(\nu + 2)}{\nu(n_i + \nu + 2)}.$$

3. If the  $i$ th subject has both censored and non-censored measurements, then

$$\begin{aligned} \widehat{\tau}_i^2 &= E_{\mathbf{y}_i^c} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \middle| \mathbf{u}_i, \mathbf{c}_i \right] = E_{\mathbf{y}_i^c} \left\{ E_{\tau_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \middle| \mathbf{y}_i \right] \middle| \mathbf{u}_i, \mathbf{c}_i \right\} \\ &= \frac{d_{n_i}(n_i^o, \nu, 2) \mathcal{T}_{n_i - n_i^o}(\mathbf{u}_i^c | \boldsymbol{\mu}_i^{c \cdot o}, \frac{\nu + n_i^o}{\nu + n_i^o + 4} \mathbf{S}_i^{cc \cdot o}, \nu + n_i^o + 4)}{(\nu + \Delta_{\mathbf{y}_i^o})^2 \mathcal{T}_{n_i - n_i^o}(\mathbf{u}_i^c | \boldsymbol{\mu}_i^{c \cdot o}, \mathbf{S}_i^{cc \cdot o}, \nu + n_i^o)}, \\ \widehat{\tau}^2\mathbf{y}_i &= E_{\mathbf{y}_i^c} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \middle| \mathbf{u}_i, \mathbf{c}_i \right] = E_{\mathbf{y}_i^c} \left\{ E_{\tau_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \middle| \mathbf{y}_i \right] \middle| \mathbf{u}_i, \mathbf{c}_i \right\} \\ &= \widehat{\tau}_i^2 E[\mathbf{W}_i^{c*}], \\ \widehat{\tau}^2\mathbf{y}_i^2 &= E_{\mathbf{y}_i^c} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \mathbf{y}_i^T \middle| \mathbf{u}_i, \mathbf{c}_i \right] = E_{\mathbf{y}_i^c} \left\{ E_{\tau_i} \left[ \left(\frac{\nu + n_i}{\nu + \Delta_{\mathbf{y}_i}}\right)^2 \mathbf{y}_i \mathbf{y}_i^T \middle| \mathbf{y}_i \right] \middle| \mathbf{u}_i, \mathbf{c}_i \right\} \\ &= \widehat{\tau}_i^2 E[\mathbf{W}_i^{c*} \mathbf{W}_i^{c*T}], \end{aligned}$$

where  $\mathbf{W}_i^{c*} \sim \mathcal{T}t_{n_i - n_i^o}(\boldsymbol{\mu}_i^{c \cdot o}, \frac{\nu + n_i^o}{\nu + n_i^o + 4} \mathbf{S}_i^{cc \cdot o}, \nu + n_i^o + 4)$ , and

$$d_{n_i}(n_i^o, \nu, 2) = (\nu + n_i)^2 \frac{\Gamma(\frac{n_i + \nu}{2}) \Gamma(\frac{n_i^o + \nu + 4}{2})}{\Gamma(\frac{n_i^o + \nu}{2}) \Gamma(\frac{n_i + \nu + 4}{2})} = \frac{(n_i + \nu)(n_i^o + \nu + 2)(n_i^o + \nu)}{(n_i + \nu + 2)}.$$

□

## Web Appendix D: Supplementary table for the simulation study.

**Table S.1.** Simulation results for comparisons of four considered models in terms of model selection and parameter estimation under the scenario of  $d = 0.50$  and  $\varrho = 0.75$  based on 100 replications.

| DOF        | Censoring | Model | Model Fitting |        |         |        | Estimation of Fixed Effects |       |              |       |              |       |              |                    |
|------------|-----------|-------|---------------|--------|---------|--------|-----------------------------|-------|--------------|-------|--------------|-------|--------------|--------------------|
|            |           |       | AIC           | (Freq) | BIC     | (Freq) | $\beta_{10}$                | (SE)  | $\beta_{11}$ | (SE)  | $\beta_{20}$ | (SE)  | $\beta_{21}$ | (SE)               |
| $\nu = 5$  | 5%        | Nc    | 1714.13       | (0)    | 1726.32 | (0)    | 1.259                       | 0.080 | 2.005        | 0.001 | -1.716       | 0.061 | 4.003        | $5 \times 10^{-4}$ |
|            |           | Tc    | 1620.86       | (100)  | 1634.27 | (100)  | 1.063                       | 0.081 | 2.002        | 0.001 | -1.879       | 0.060 | 3.997        | $5 \times 10^{-4}$ |
|            |           | N     | 1726.48       | (0)    | 1738.67 | (0)    | 1.282                       | 0.330 | 1.968        | 0.038 | -1.677       | 0.281 | 3.961        | 0.028              |
|            |           | T     | 1651.39       | (0)    | 1663.58 | (0)    | 1.197                       | 0.288 | 1.975        | 0.032 | -1.716       | 0.252 | 3.968        | 0.023              |
|            | 10%       | Nc    | 1628.79       | (0)    | 1640.98 | (1)    | 1.204                       | 0.073 | 2.015        | 0.001 | -1.747       | 0.058 | 4.010        | $4 \times 10^{-4}$ |
|            |           | Tc    | 1541.42       | (100)  | 1554.83 | (99)   | 1.016                       | 0.071 | 2.011        | 0.001 | -1.858       | 0.056 | 3.998        | $4 \times 10^{-4}$ |
|            |           | N     | 1759.11       | (0)    | 1771.30 | (0)    | 1.798                       | 0.316 | 1.899        | 0.038 | -1.071       | 0.279 | 3.878        | 0.030              |
|            |           | T     | 1698.21       | (0)    | 1710.40 | (0)    | 1.713                       | 0.289 | 1.909        | 0.032 | -1.053       | 0.263 | 3.884        | 0.026              |
|            | 20%       | Nc    | 1491.82       | (0)    | 1504.01 | (0)    | 1.215                       | 0.061 | 2.011        | 0.001 | -1.793       | 0.055 | 4.016        | $4 \times 10^{-4}$ |
|            |           | Tc    | 1405.51       | (100)  | 1418.92 | (100)  | 1.068                       | 0.062 | 1.997        | 0.001 | -1.918       | 0.051 | 3.998        | $3 \times 10^{-4}$ |
|            |           | N     | 1843.98       | (0)    | 1856.17 | (0)    | 3.262                       | 0.301 | 1.708        | 0.039 | 0.156        | 0.292 | 3.713        | 0.037              |
|            |           | T     | 1806.64       | (0)    | 1818.83 | (0)    | 3.208                       | 0.302 | 1.714        | 0.036 | 0.170        | 0.298 | 3.718        | 0.035              |
| $\nu = 50$ | 5%        | N     | 1486.05       | (88)   | 1498.24 | (93)   | 0.952                       | 0.077 | 2.005        | 0.001 | -1.993       | 0.056 | 3.997        | $5 \times 10^{-4}$ |
|            |           | Tc    | 1487.57       | (12)   | 1500.98 | (7)    | 0.954                       | 0.077 | 2.005        | 0.001 | -1.992       | 0.055 | 3.997        | $5 \times 10^{-4}$ |
|            |           | N     | 1516.06       | (0)    | 1528.25 | (0)    | 1.146                       | 0.271 | 1.979        | 0.031 | -1.770       | 0.230 | 3.967        | 0.022              |
|            |           | T     | 1514.78       | (0)    | 1526.97 | (0)    | 1.142                       | 0.277 | 1.979        | 0.031 | -1.766       | 0.243 | 3.967        | 0.022              |
|            | 10%       | Nc    | 1407.98       | (89)   | 1420.17 | (97)   | 0.868                       | 0.070 | 2.017        | 0.001 | -2.014       | 0.054 | 4.001        | $4 \times 10^{-4}$ |
|            |           | Tc    | 1409.52       | (11)   | 1422.93 | (3)    | 0.873                       | 0.071 | 2.016        | 0.001 | -2.012       | 0.053 | 4.000        | $4 \times 10^{-4}$ |
|            |           | N     | 1572.64       | (0)    | 1584.83 | (0)    | 1.650                       | 0.259 | 1.911        | 0.031 | -1.164       | 0.231 | 3.884        | 0.025              |
|            |           | T     | 1571.75       | (0)    | 1583.93 | (0)    | 1.652                       | 0.277 | 1.912        | 0.030 | -1.147       | 0.254 | 3.885        | 0.025              |
|            | 20%       | Nc    | 1275.91       | (93)   | 1288.10 | (98)   | 0.987                       | 0.058 | 2.000        | 0.001 | -2.001       | 0.050 | 3.999        | $4 \times 10^{-4}$ |
|            |           | Tc    | 1277.56       | (7)    | 1290.96 | (2)    | 0.989                       | 0.063 | 2.000        | 0.001 | -2.002       | 0.050 | 3.999        | $3 \times 10^{-4}$ |
|            |           | N     | 1694.51       | (0)    | 1706.70 | (0)    | 3.133                       | 0.252 | 1.716        | 0.033 | 0.090        | 0.251 | 3.716        | 0.033              |
|            |           | T     | 1694.33       | (0)    | 1706.52 | (0)    | 3.137                       | 0.288 | 1.717        | 0.033 | 0.101        | 0.287 | 3.716        | 0.033              |

Nc: MLMMC; Tc: MlMMC; N: MLMM; T: MlMM